

Copulas ? What copulas ?

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Multivariate linear correlations

- Standard tool in risk management/portfolio optimisation: the covariance matrix $R_{ij} = \langle r_i r_j \rangle$

- Find the portfolio with maximum expected return for a given risk or equivalently, minimum risk for a given return (\mathcal{G})

- In matrix notation:

$$\mathbf{w} = \mathcal{G} \frac{\mathbf{R}^{-1} \mathbf{g}}{\mathbf{g}^T \mathbf{R}^{-1} \mathbf{g}}$$

where all gains are measured with respect to the risk-free rate and $\sigma_i = 1$ (absorbed in g_i).

- More explicitly:

$$\mathbf{w} \propto \sum_{\alpha} \lambda_{\alpha}^{-1} (\Psi_{\alpha} \cdot \mathbf{g}) \Psi_{\alpha} = \mathbf{g} + \sum_{\alpha} (\lambda_{\alpha}^{-1} - 1) (\Psi_{\alpha} \cdot \mathbf{g}) \Psi_{\alpha}$$

Multivariate non-linear correlations

- Many situations in finance in fact require knowledge of **higher order correlations**
 - **Gamma-risk of option portfolios:** $\langle r_i^2 r_j^2 \rangle - \langle r_i^2 \rangle \langle r_j^2 \rangle$
 - **Stress test of complex portfolios:** correlations in extreme market conditions
 - **Correlated default probabilities** – Credit Derivatives (CDOs, basket of CDSs)
 - “The Formula That Killed Wall Street” (Felix Salmon)

Different correlation coefficients

- Correlation coefficient: $\rho_{ij} = \text{COV}(r_i, r_j) / \sqrt{V(r_i)V(r_j)}$

- Correlation of squares or absolute values:

$$\rho_{ij}^{(2)} = \frac{\text{COV}(r_i^2, r_j^2)}{\sqrt{V(r_i^2)V(r_j^2)}} \quad \rho_{ij}^{(a)} = \frac{\text{COV}(|r_i|, |r_j|)}{\sqrt{V(|r_i|)V(|r_j|)}}$$

- Tail correlation:

$$\tau_{ij}^{UU}(p) = \frac{1}{p} \text{Prob.} [r_i > \mathcal{P}_{>,i}^{-1}(p) \cap r_j > \mathcal{P}_{>,j}^{-1}(p)]$$

(Similar defs. for $\tau^{LL}, \tau^{UL}, \tau^{LU}$)

Copulas

- **Sklar's theorem:** any multivariate distribution can be “factorized” into
 - its marginals $\mathcal{P}_i \rightarrow u_i = \mathcal{P}_i(r_i)$ are $U[0, 1]$
 - a “copula”, that describes the correlation structure between N $U[0, 1]$ standardized random variables: $c(u_1, u_2, \dots, u_N)$
- All correlations, linear and non linear, can be computed from the copula and the marginals
- For bivariate distributions:

$$C_{ij}(u, v) = \mathbb{P} \left[\mathcal{P}_{<,i}(X_i) \leq u \text{ and } \mathcal{P}_{<,j}(X_j) \leq v \right]$$

Copulas – Examples

- **Examples:** ($N = 2$)
 - **The Gaussian copula:** r_1, r_2 bivariate Gaussian \rightarrow defines the Gaussian copula $c_G(u, v | \rho)$
 - **The Student copula:** r_1, r_2 bivariate Student with tail $\nu \rightarrow$ defines the Student copula $c_S(u, v | \rho, \nu)$
 - **Archimedean copulas:** $\phi(u) : [0, 1] \rightarrow [0, 1]$, $\phi(1) = 0$, ϕ^{-1} decreasing, completely monotone

$$C_A(u, v) = \phi^{-1} [\phi(u) + \phi(v)]$$

Ex: **Frank copulas**, $\phi(u) = \ln[e^\theta - 1] - \ln[e^{\theta u} - 1]$;

Gumbel copulas, $\phi(u) = (-\ln u)^\theta$, $\theta < 1$.

The Copula red-herring

- Sklar's theorem: a nearly empty shell – almost any $c(u_1, u_2, \dots, u_N)$ with required properties is allowed.
- The usual financial mathematics syndrom: choose a class of copulas – sometimes absurd – with convenient mathematical properties and brute force calibrate to data
- If something fits it can't be bad (??) Statistical tests are not enough – intuition & plausible interpretation are required
- *But he does not wear any clothes!* – see related comments by Thomas Mikosch

The Copula red-herring

- **Example 1:** why on earth choose the Gaussian copula to describe correlation between (positive) default times???
- **Example 2:** Archimedean copulas: take two $U[0, 1]$ random variables s, w . Set $t = K^{-1}(w)$ with $K(t) = t - \phi(t)/\phi'(t)$.

$$u = \phi^{-1} [s\phi(t)]; \quad v = \phi^{-1} [(1 - s)\phi(t)]; \quad \longrightarrow r_1, r_2$$

Financial interpretation ???

- Models should reflect some plausible underlying structure or mechanism

Copulas ? What copulas ?

- Aim of this work
 - Develop intuition around copulas
 - Identify empirical stylized facts about multivariate correlations that copulas should reproduce
 - Discuss “self-copulas” as a tool to study empirical temporal dependences
 - Propose an intuitively motivated, versatile model to generate a wide class of non-linear correlations

Copulas

- **Restricted information on copula:** diagonal $C(p, p)$ and anti-diagonal $C(p, 1 - p)$. Note: $C(\frac{1}{2}, \frac{1}{2})$ is the probability that both variables are simultaneously below their medians

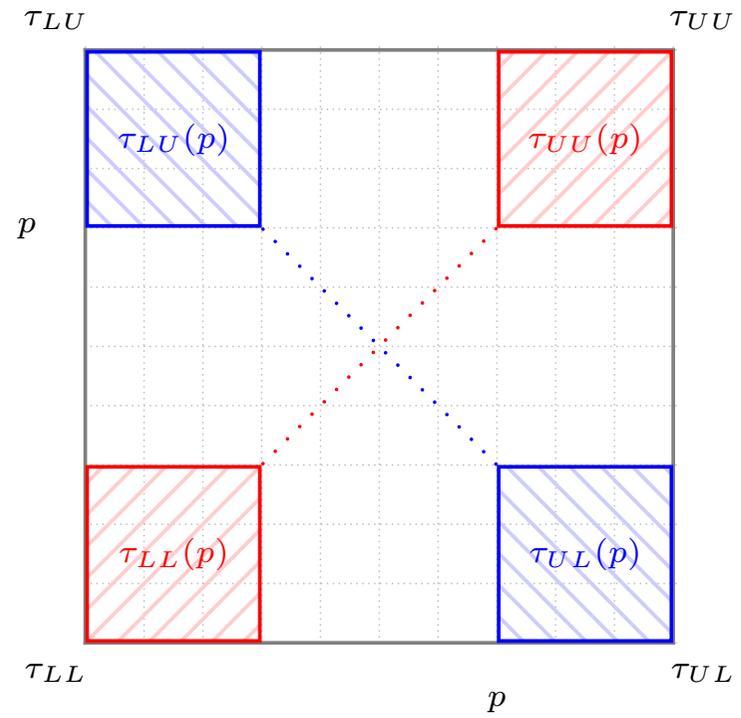
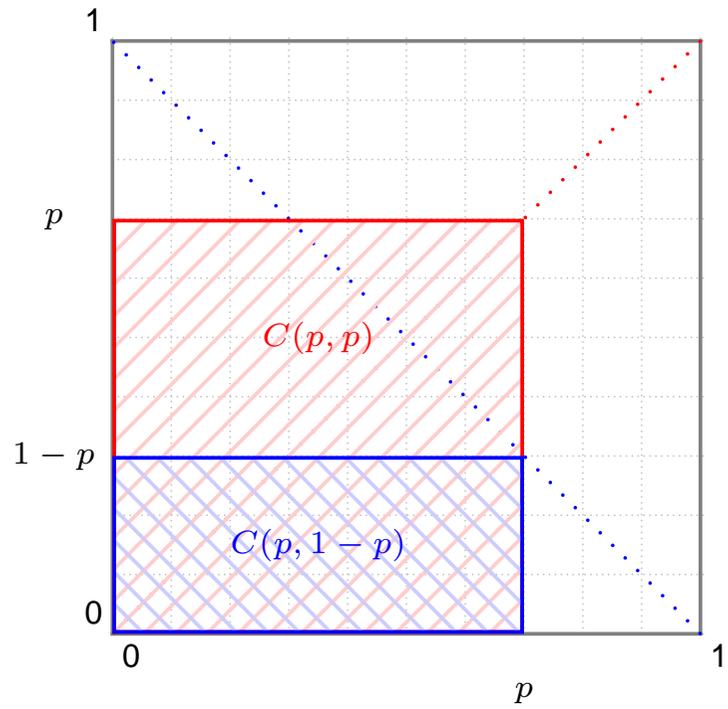
- **Tail dependence:**

$$\tau^{UU}(p) = \frac{1 - 2p + C(p, p)}{1 - p}, \quad \text{etc.}$$

- **Relative difference** with respect to independence or to Gaussian:

$$\frac{C(p, p) - p^2}{p(1 - p)} = \tau^{UU}(p) + \tau^{LL}(1 - p) - 1, \quad \text{or} \quad \frac{C(p, p) - C_G(p, p)}{p(1 - p)}$$

Copulas



Student Copulas

- **Intuition:** $r_1 = \sigma\epsilon_1$, $r_2 = \sigma\epsilon_2$ with:
 - $\epsilon_{1,2}$ bivariate Gaussian with correlation ρ
 - σ is a **common random volatility** with distribution $P(\sigma) = \mathcal{N}e^{-\sigma_0^2/\sigma^2} / \sigma^{1+\nu}$
- **The univariate distributions** of $r_{1,2}$ are Student with a power law tail exponent ν ($\in [3, 5]$ for daily data)
- **The multivariate Student** is a model of correlated Gaussian variables with a **common random volatility**:

$$r_i = \sigma\epsilon_i \quad \rho_{ij} = \text{COV}(\epsilon_i, \epsilon_j)$$

Student Copulas

- In this model, all higher-order correlations can be expressed in terms of ρ

- Explicit formulas: ($f_n = \langle \sigma^{2n} \rangle / \langle \sigma^n \rangle^2$)

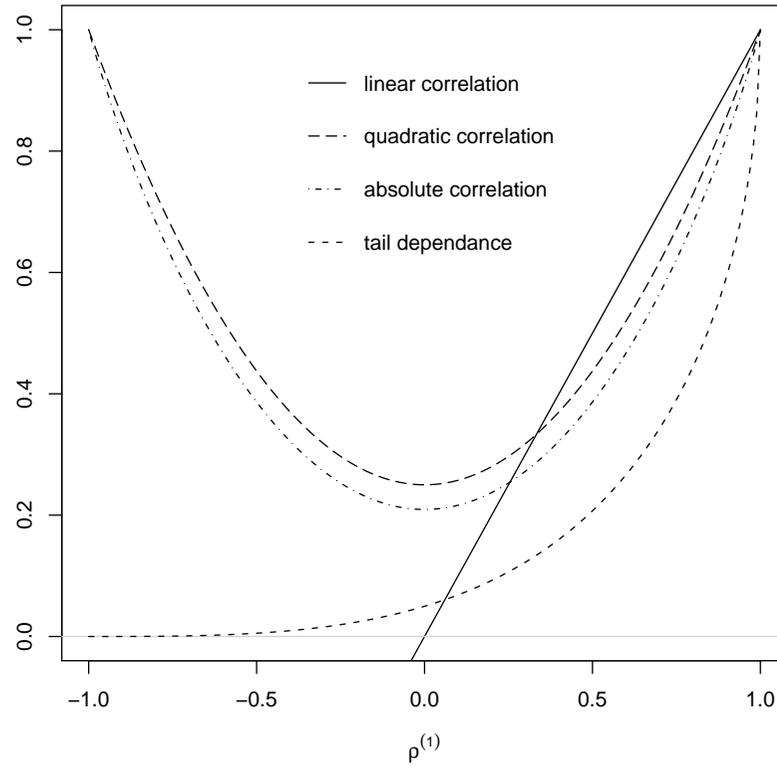
$$\rho^{(2)} = \frac{f_2(1 + 2\rho^2) - 1}{3f_2 - 1}; \quad \rho^a = \frac{f_1(\sqrt{1 - \rho^2} + \rho \arcsin \rho) - 1}{\frac{\pi}{2}f_1 - 1}.$$

- The tail correlations τ have a finite limit when $p \rightarrow 0$ because of the common volatility

- The central point of the copula:

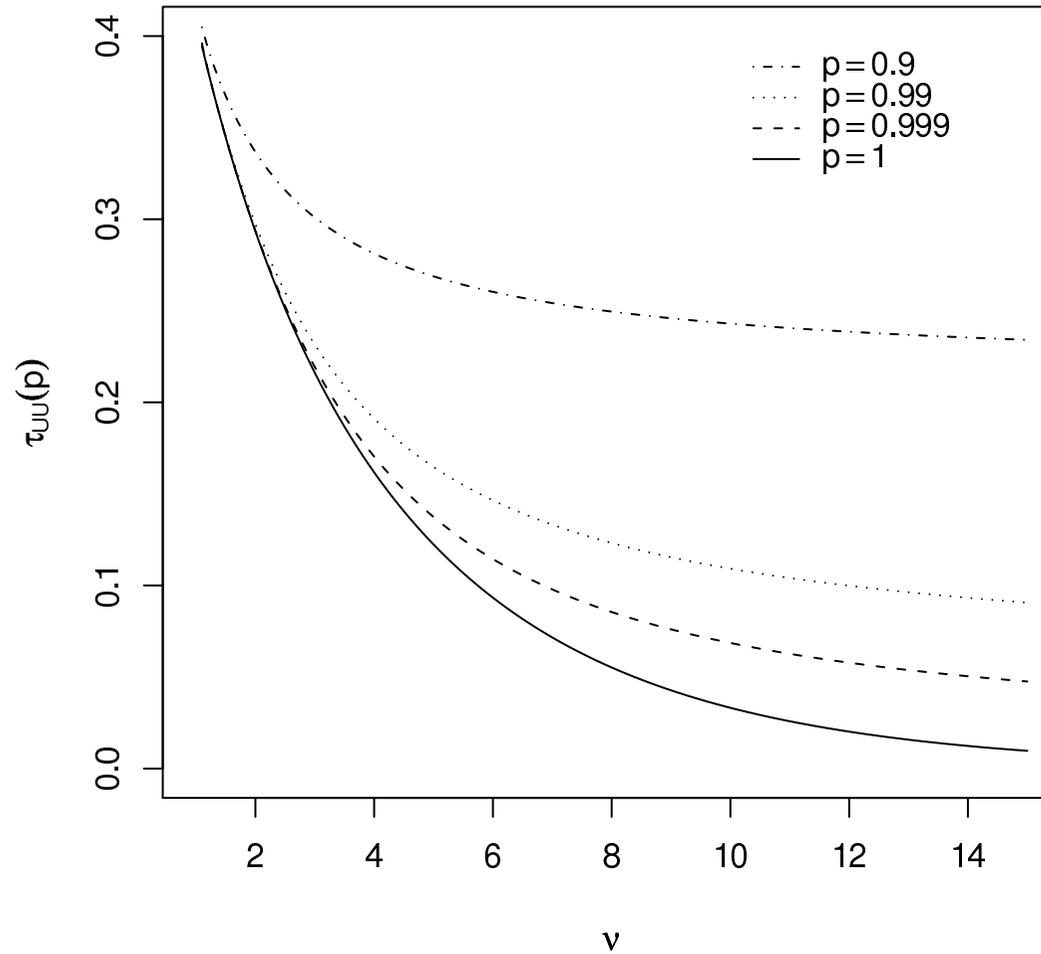
$$C\left(\frac{1}{2}, \frac{1}{2}\right) = \frac{1}{4} + \frac{1}{2\pi} \arcsin \rho$$

Student Copulas



$$\nu = 5$$

Student Copulas



$\rho = 0.3$ – Note: corrections are of order $(1 - p)^{2/\nu}$

Elliptic Copulas

- A straight-forward generalisation: elliptic copulas

$$r_1 = \sigma\epsilon_1, r_2 = \sigma\epsilon_2, P(\sigma) \text{ arbitrary}$$

- The above formulas remain valid for arbitrary $P(\sigma)$ in particular:

$$C\left(\frac{1}{2}, \frac{1}{2}\right) = \frac{1}{4} + \frac{1}{2\pi} \arcsin \rho$$

- The tail correlations τ have a finite limit whenever $P(\sigma)$ decays as a power-law

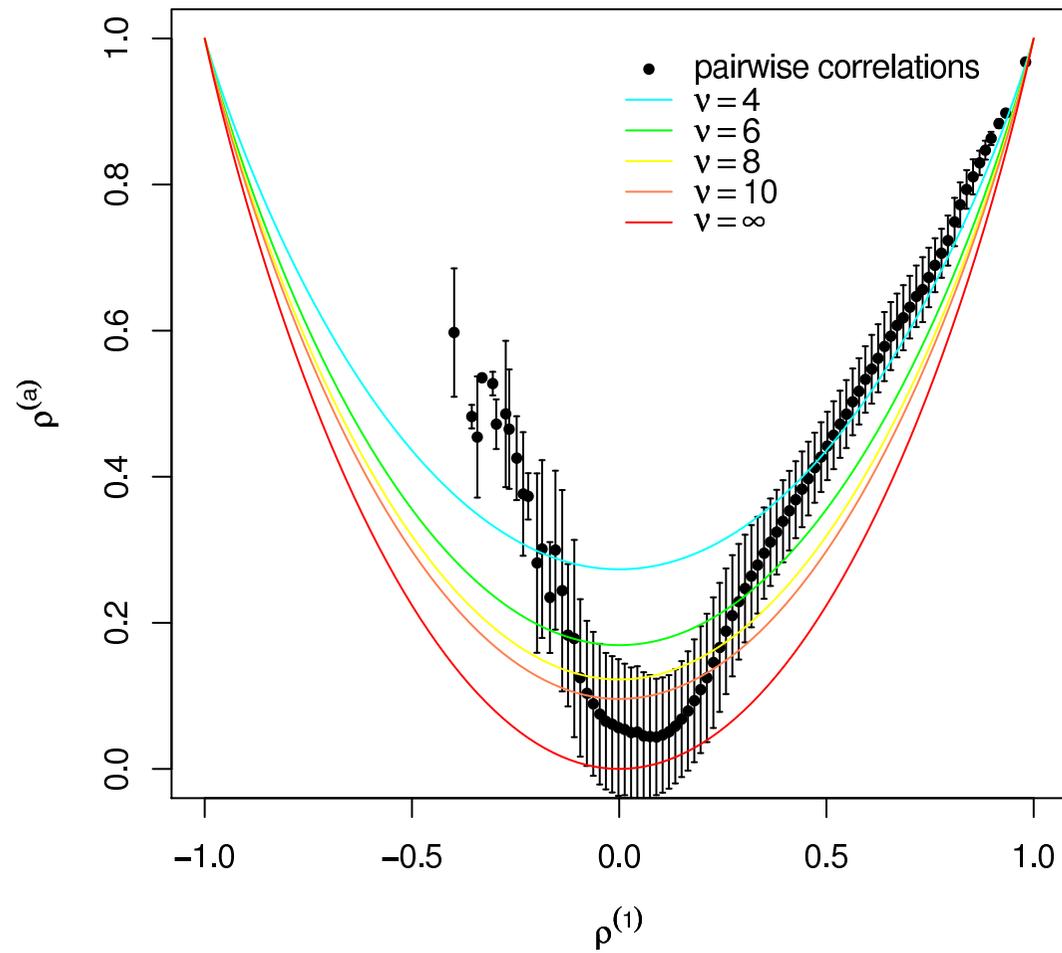
- A relevant example: the log-normal model $\sigma = \sigma_0 e^\xi$, $\xi = N(0, \lambda^2)$ – very similar to Student with $\nu \sim \lambda^{-2}$ *

*Although the true asymptotic value of $\tau(p=0)$ is zero.

Student Copulas and empirical data

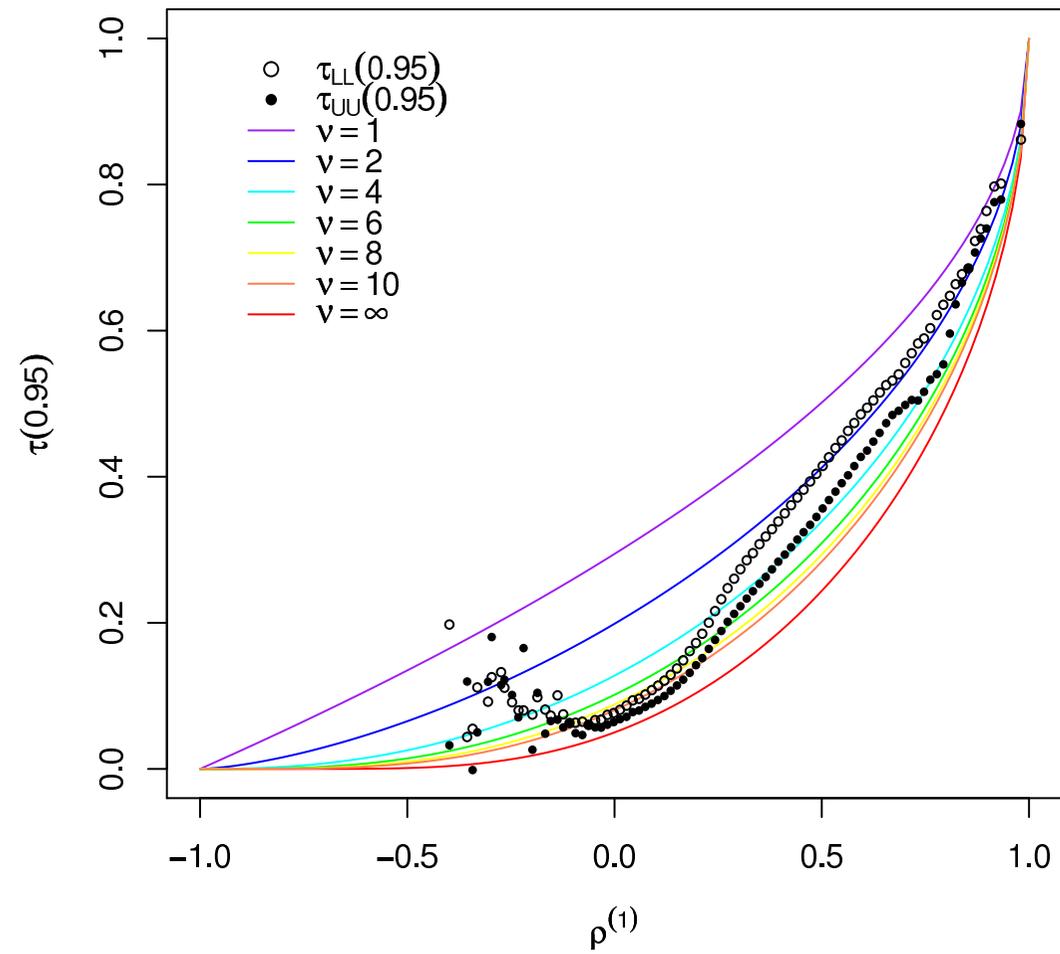
- The empirical curves $\rho^a(\rho)$ or $\rho^{(2)}(\rho)$ **cross the set of Student predictions**, as if “more Gaussian” for small ρ 's
- **Same with tail correlation coefficients** (+ some level of assymetry)
- $C(\frac{1}{2}, \frac{1}{2})$ **systematically different** from Elliptic prediction
 $= \frac{1}{4} + \frac{1}{2\pi} \arcsin \rho$ – in particular $C(\frac{1}{2}, \frac{1}{2} | \rho = 0) > \frac{1}{4}$
- $C(p, p) - C_G(p, p)$ **incompatible with a Student model**: concave for $\rho < 0.25$ becoming convex for $\rho > 0.25$
- **To be sure: Archimedean copulas are even worse !**

Absolute correlation



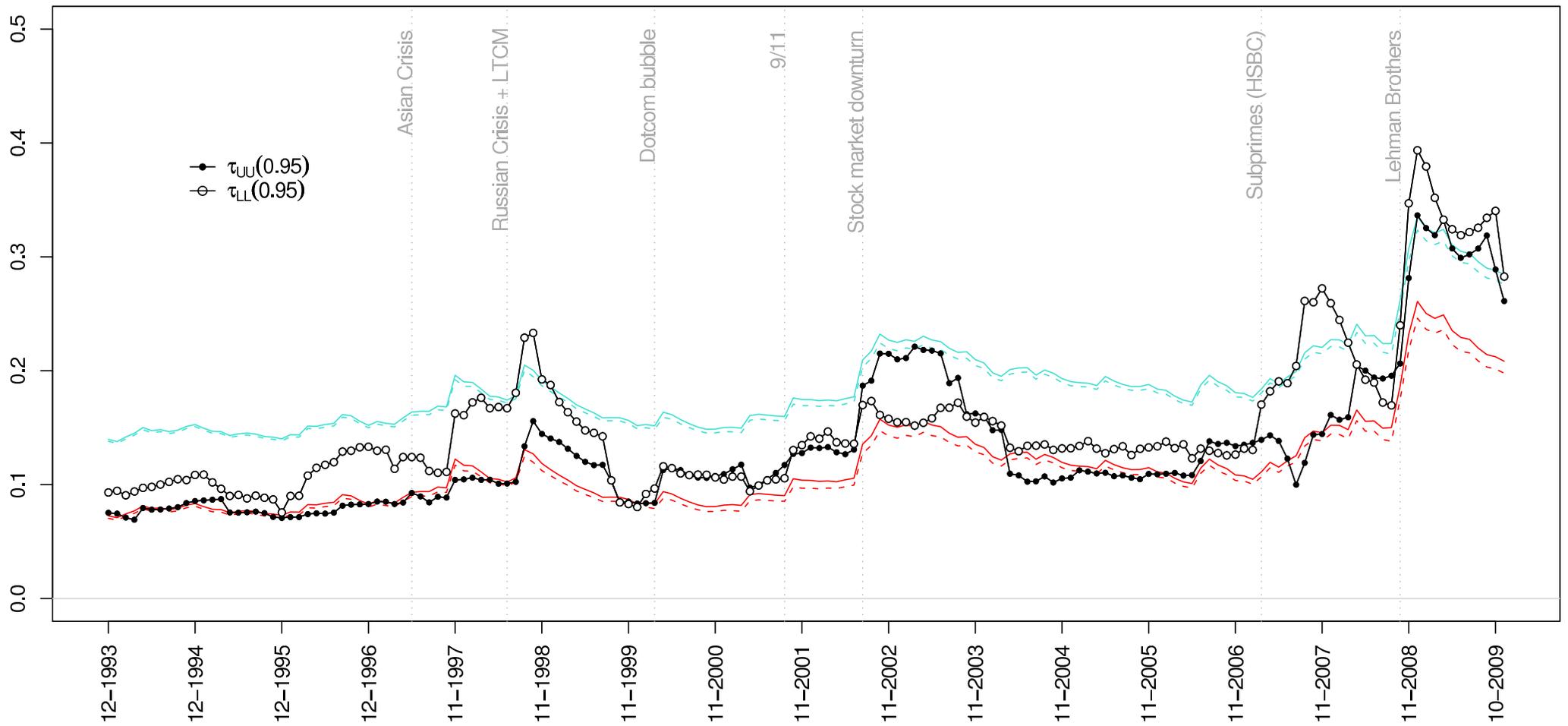
2005-2009

Tail correlation

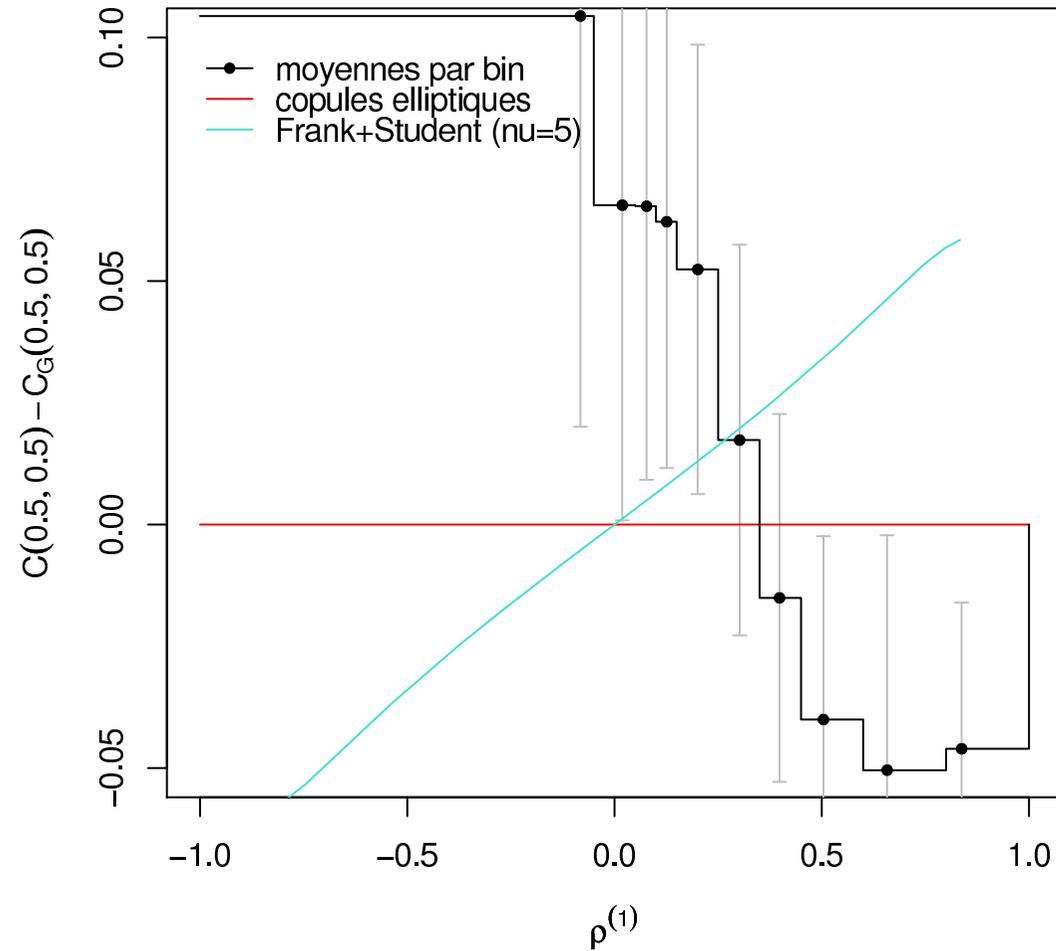


2005-2009

Tail correlation: time series

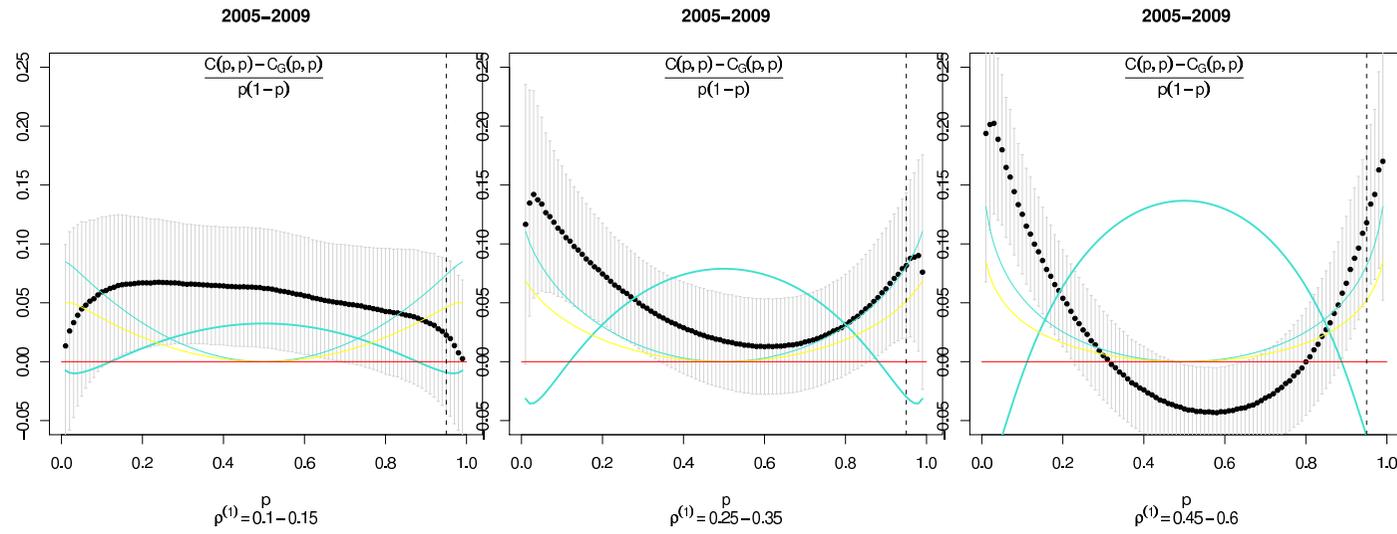


Centre point vs ρ



Difference between empirical results and Student (Frank)
prediction for $C(\frac{1}{2}, \frac{1}{2})$

Diagonal



$$\rho = 0.1, 0.3, 0.5$$

Student Copulas: Conclusion

- Student (or even elliptic) copulas are not sufficient to describe the multivariate distribution of stocks!
- **Obvious intuitive reason:** one expects more than one volatility factor to affect stocks
- How to describe an **entangled correlation** between returns and volatilities?
- In particular, any model such that $r_i = \sigma_i \epsilon_i$ with correlated random σ 's leads to $C(\frac{1}{2}, \frac{1}{2}) \equiv \frac{1}{4}$ for $\rho = 0$!

Constructing a realistic copula model

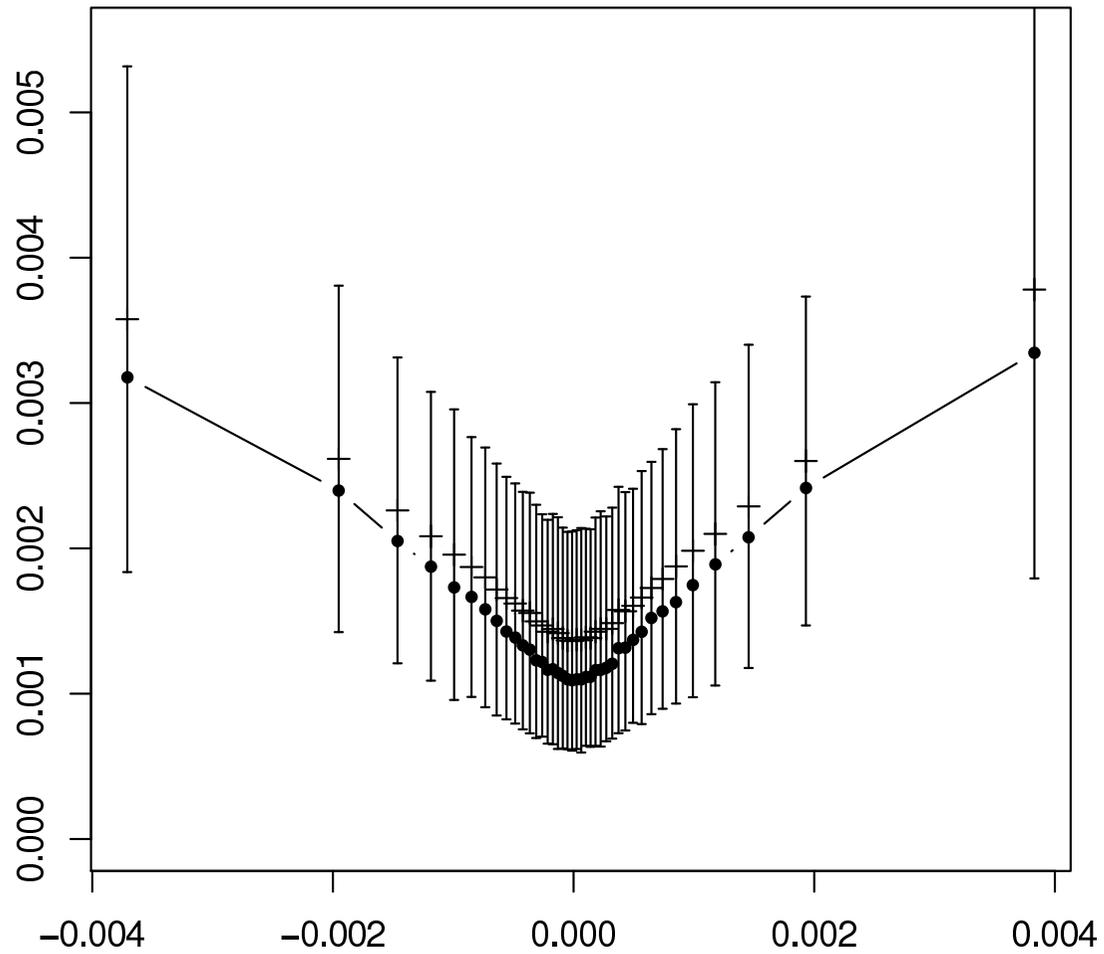
- How do we go about now (for stocks) ?
 - a) stocks are sensitive to “factors”
 - b) factors are hierarchical, in the sense that the vol of the market influences that of sectors, which in turn influence that of more idiosyncratic factors

- Empirical fact: within a one-factor model,

$$r_i = \beta_i \varepsilon_0 + \varepsilon_i$$

volatility of residuals increases with that of the market ε_0

“Entangled” volatilities



with Romain Allez

Constructing a realistic copula model

- An entangled one-factor model

$$r_i = \beta_i \sigma_0 e^{\xi_0} \varepsilon_0 + \sigma_1 e^{\alpha \xi_0 + \xi_i} \varepsilon_i$$

with $\xi_0 \sim N(0, s_0^2)$, $\xi_i \sim N(0, s_1^2)$, IID,

- The volatility of the idiosyncratic factor is clearly affected by that of **the market mode**
- **Kurtosis** of the market factor and of the idiosyncratic factor:

$$\kappa_0 = e^{4s_0^2} \left[e^{4s_0^2} - 1 \right]; \quad \kappa_1 = e^{4(\alpha^2 s_0^2 + s_1^2)} \left[e^{4(\alpha^2 s_0^2 + s_1^2)} - 1 \right]$$

Constructing a realistic copula model

- An interesting remark: take two stocks with opposite exposure to the second factor

$$r_{\pm} = \sigma_0 e^{\xi_0} \varepsilon_0 \pm \sigma_1 e^{\alpha \xi_0 + \xi_1} \varepsilon_1$$

- Choose parameters such that volatilities are equal

$$\sigma_0 e^{s_0^2} = \sigma_1 e^{\alpha^2 s_0^2 + s_1^2}$$

such that $\text{cov}(r_+, r_-) = 0$

- Then:

$$C\left(\frac{1}{2}, \frac{1}{2} \mid \rho = 0\right) \approx \frac{1}{4} \left(1 + \frac{\kappa_1 - \kappa_0}{6\pi}\right)$$

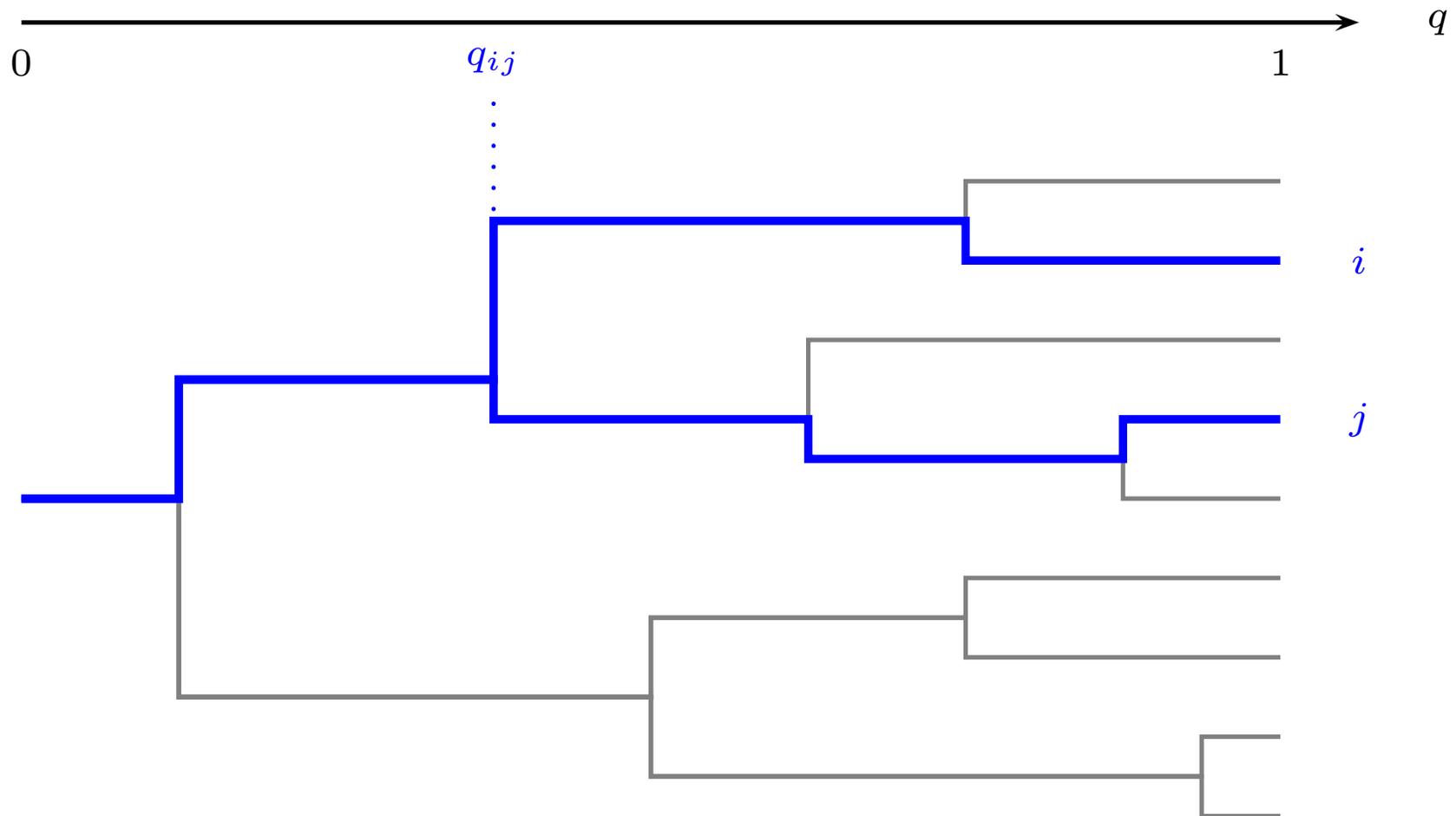
A hierarchical tree model

- Construct a tree such that the trunk is the market factor, and each link is a factor with entangled vol.
- The return of stock i is constructed by following a path C_i along the tree from trunk to leaves

$$r_i = \int_{C_i, q \in [0,1]} \beta_i(q) \sigma(q) d\varepsilon(q) \times \exp \left[\int_{C_i, q' \in [0,q]} \alpha(q, q') d\xi(q') \right]$$

- **Parameters:** Branching ratio of the tree $b(q)$, volatility function $\sigma(q)$, intrication function $\alpha(q, q')$

A hierarchical tree model



A hierarchical tree model

- Calibration on data: work in progress...
- Find simple, systematic ways to calibrate such a huge model \Rightarrow stability of R_{ij} ??
- Preliminary simulation results for reasonable choices: the model is able to reproduce all the empirical facts reported above, including $C(1/2, 1/2) > 1/4$ and the change of concavity of

$$\frac{C(p, p) - C_G(p, p)}{p(1 - p)}$$

as ρ increases

Self-copulas

- One can also define the copula between a variable and itself, lagged:

$$C_\tau(u, v) = \mathbb{P} \left[\mathcal{P}_<(X_t) \leq u \text{ and } \mathcal{P}_{<,j}(X_{t+\tau}) \leq v \right]$$

- Example: log-normal copula

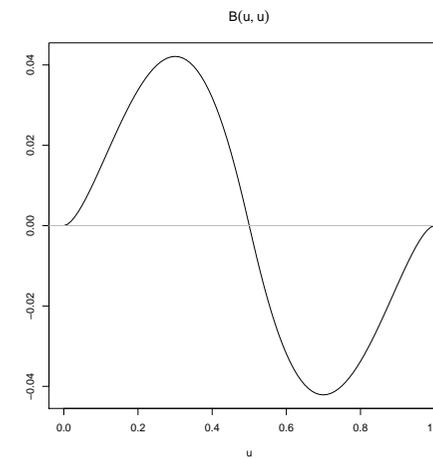
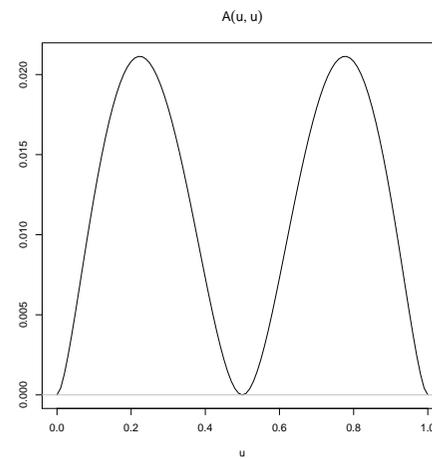
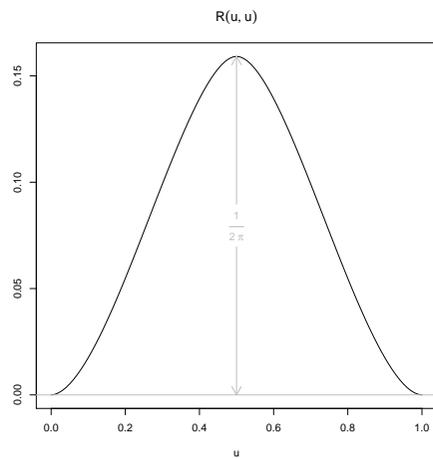
$$X_t = e^{\omega t} \xi_t$$

with correlations between ξ 's (linear), ω 's (vol) and $\omega\xi$ (leverage)

- In the limit of weak correlations:

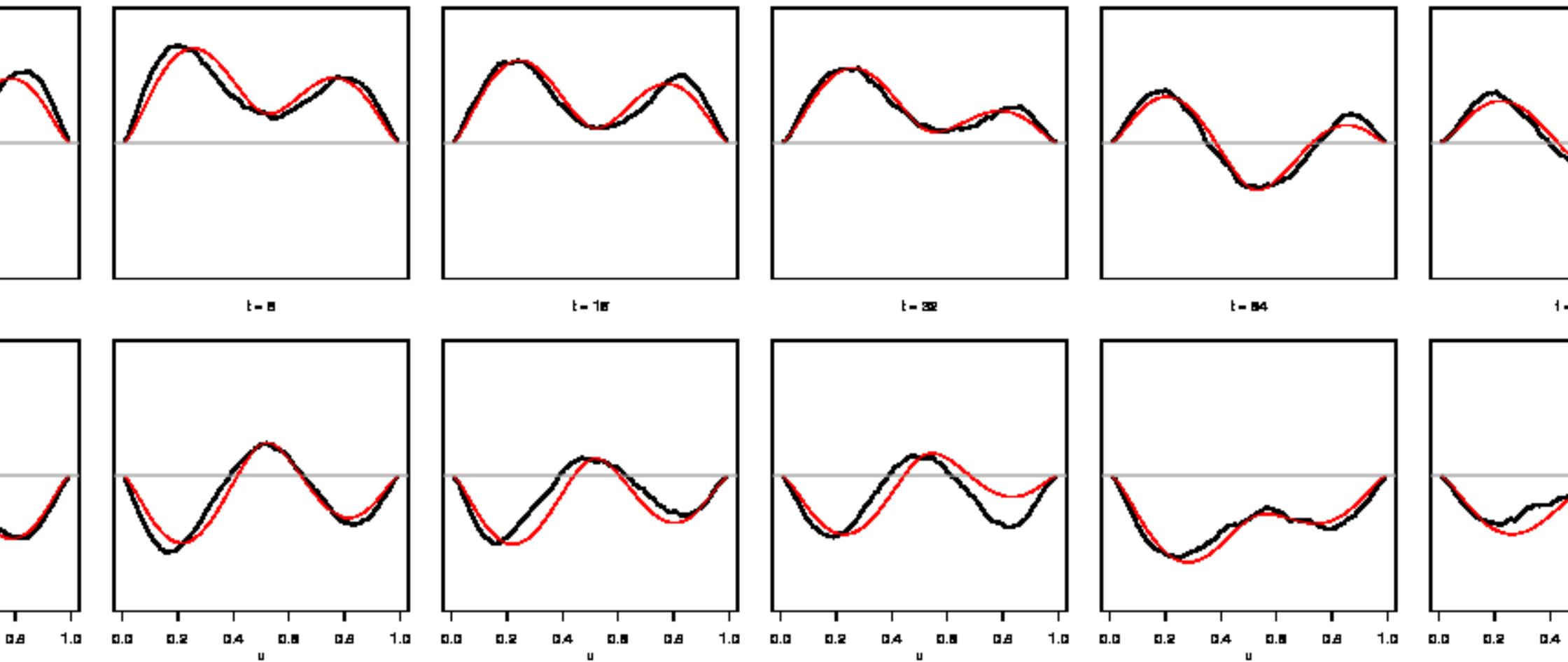
$$C_t(u, v) - uv \approx \rho R(u, v) + \alpha A(u, v) - \beta B(u, v)$$

Three corrections to independence

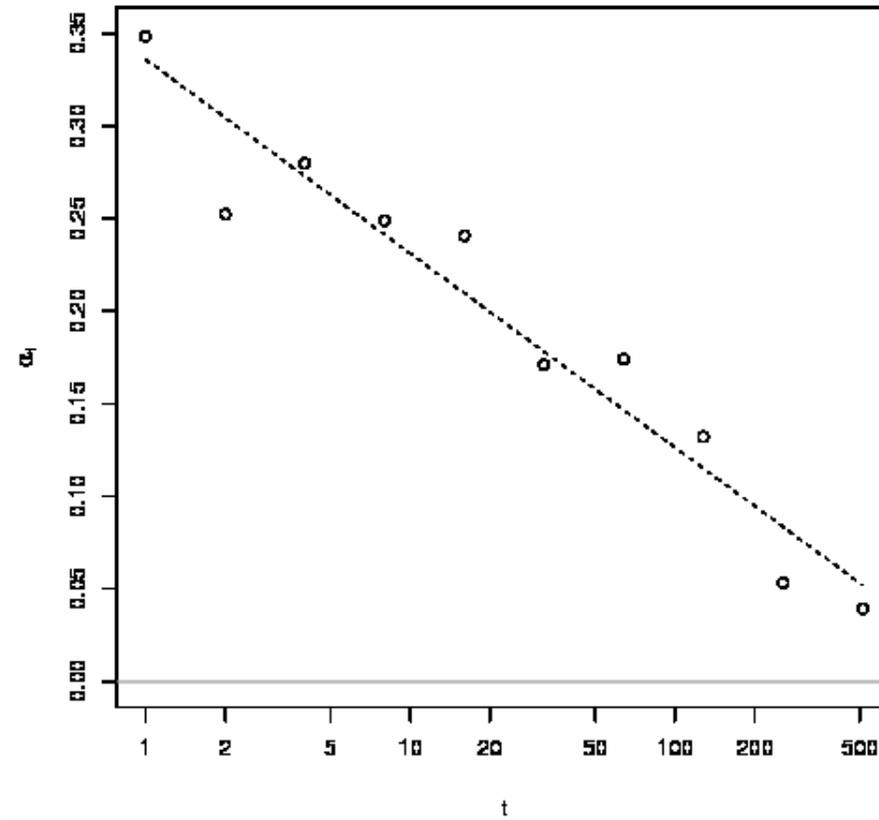


ρ, α, β

Empirical self copulas



Long range (multifractal) memory



Self-copulas

- **A direct application:** GoF tests (Kolmogorov-Smirnov/Cramer-von Mises) for dependent variables
- The relevant quantity is $\sum_t (C_t(u, v) + C_{-t}(u, v) - 2uv)$
- The test is dependent on the self-copula
- \Rightarrow Significant decrease of the effective number of independent variables

Conclusion – Open problems

- GoF tests for two-dimensional copulas: max of “Brownian sheets” (some progress with Rémy)
- Structural model: requires analytical progress (possible thanks to the tree structure) and numerical simulations
- Extension to account for U/L asymmetry
- Extension to describe defaults and time to defaults – move away from silly models and introduce some underlying *structure*